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Genetic Evaluation with a Multiple Lactation Test Day Model for SCS and Production Traits

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ABSTRACT

A multiple lactation test day model was applied to predict genetic merit for somatic cell score and protein yield in Canadian Holstein cattle. The model for test day genetic evaluation included a fixed herd-test-date effect, fixed regressions on functions of days in milk, random permanent environmental effects within lactation, random animal genetic effects, and residual effects. Records from the first three lactations were used and treated as different traits. Procedures for this model were developed for national genetic evaluation for somatic cell score and were found to be practical even for a much larger dataset from western Germany. Use of starting values from the previous genetic evaluation run reduced the number of rounds necessary to reach convergence. Test day models were compared to several single trait models based on lactation averages in terms of ranking of animals. Differences between EBV from the test-day model and EBV from a lactation average repeatability model were small for bulls with many daughters, but differences with EBV from a single trait first lactation average model were large. Differences were smaller for SCS than for protein yield.

OBJECTIVES

The objectives of this study were:

a) to develop procedures for national genetic evaluation based on multiple lactation test day models and to apply them to field data (from Canada and Germany)

b) to compare test-day models to lactation average models for SCC and for protein yield. Two different test day models were compared to models for genetic evaluation based on lactation measures using the data set from Canada (data set 1). Results on computational aspects will be presented for the larger German data set (data set 2).

MATERIALS AND METHODS

Data Data consisted of test day records for SCC and production traits from data records processing centres in Canada and Germany, which were analyzed separately. For Canada, data included records from test days from 1987 until fall of 1994. For Germany, the time span was from 1990 to summer 1995. For each test day, SCC was transformed to SCS to achieve normality and homogeneity of variances.

Table 1 shows numbers of records and numbers of herd-test-date levels and cows. Following edits, 5,505,672 records on 392,487 cows remained for Canada and 22,617,153 records on 1,487,279 cows remained in the German data set. Pedigree was completed for cows with identification of dam and maternal grandsire from national pedigree files. Pedigrees for bulls with daughter records or granddaughter records were completed for several generations. Unknown parents were assigned to phantom parent groups, grouped by birth year of offspring (5 years per interval).

Models

For genetic evaluation of test-day observations, a multiple trait test day model with repeated

observations within each lactation was used. Table 2 displays (co)variance matrices of additive genetic (G₀), permanent environmental (PE_0), and residual (R_0) effects, which were estimated using Gibbs sampling procedures on subsets of the analysed data sets. Material and methods for the Canadian dataset are described in a paper by Reents et al. (1995a). Variance components for Germany were estimated applying the same methodology using a dataset of 26,216 German Holstein cows with 362,478 test day records from lactations 1 to 3.

Two models were applied for analysis of test-day records. The first model (TD1) was:

 $HTD_{im} + A_{im} + P_{im} + AS_{km} + b_{km1}(D/c) + b_{km2}(D/c)^2$ y,

$$h_{\rm h} = \frac{1}{2} \ln (c/D) + b_{\rm terral} [\ln(c/D)]^2 + e_{\rm ijkmn}$$

where y_{ijkmn} is the nth test day observation of the jth cow in parity m; HTD_{im} is a fixed herd-test-date effect; A_{jm} is an animal additive genetic effect (random), P_{jm} is a within lactation permanent environmental effect to account for common environmental effects associated with all test-day records of the jth cow in lactation m (random); AS_{km} is an age-season subclass mean effect in parity m; b_{km1} and b_{km2} are regression coefficients on the linear and quadratic effects of D/c, where D is days in milk and c=381; b_{km3} and b_{km4} are regression coefficients on the linear and quadratic effects of ln(c/D); eijkmn is a random residual effect. Regression coefficients were estimated within 24 parity*age of calving*season groups.

Test-day model TD2 was very similar to model TD1, except contemporary groups for second and third lactation records from a specific herd-test-date were combined into a common herd-test-day class to increase the size of subcells.

Three genetic evaluation models based on lactation average SCS or 305-d protein yield were used for comparison. Model REP1 was a single trait repeatability model:

 $y_{ijkl} = HYS_i + PA_j + P_k + A_k + e_{ijkl}$

where y_{ijkl} is the lth lactation average of the kth cow, HYS_i is a fixed herd-year-season effect, PA_j is a fixed parity by age effect, P_k is a random permanent environmental effect across lactations of cow k, A_k is a random additive genetic effect of animal k, and e_{ijkl} is a random residual effect. Residual variances were scaled according to the number of samples included in the lactation average, based on appropriate weighting factors. Only averages based on at least two test day records were included. To compute lactation average SCS, test day records were adjusted for stage of lactation by additive pre-adjustment factors. Variance components used were based on a heritability of .11 and a repeatability of .27 for SCS and a heritability of .33 and a repeatability of .5 for protein yield, as currently used in routine evaluation in Canada.

Model REP2 was similar to REP1, but only lactation yields based on at least five test-day observations were used. All observations received the same weight in the evaluation. Variance ratios were the same as used in model REP1.

Model ST was similar to REP2, except that only first lactation data were included. Only lactation yields based on at least five test-day results were used. Pre-adjustment for stage of lactation was done with the same factors as for models REP1 and REP2. A heritability of .11 for SCS and of .33 for protein yield was used.

Models TD1 and TD2 provide separate EBV for SCS in the first three lactations. EBV can be combined into an overall EBV for SCS based on their economic values, with index weights

derived based on Schneeberger et al. (1991) as: $\mathbf{b} = \mathbf{G}_0^{-1} \mathbf{G}_T \mathbf{v}$, where G_T is a matrix of genetic (co)variances between SCS in the first 3 lactations and traits in the breeding goal and v is a vector of economic values of traits in the breeding goal. The breeding goal and economic values used were as developed by Kolstad and Dekkers (1994, Unpublished) and included four traits: clinical mastitis and subclinical mastitis (SCS) in first and second lactations. Economic values for second lactation traits accounted for the expression of traits in all later lactations, assuming a genetic correlation of one. Resulting index weights, standardized to sum to one, were .26, .66, and .08 for EBV for SCS in lactations 1, 2, and 3. For protein yield equal weights were put on each lactation EBV.

Computing strategies used for iterative solution of large scale test day animal models were described in detail by Reents et al. (1995b). The important key for processing a large number of test day records is the use of efficient input and output routines of C-HP (fread) in fortran programs. This allowed solving for each effect in the model by iteration on data (Schaeffer and Kennedy, 1986). Considerable amount of random access memory was saved by using an implicit representation of the mixed model equations for a multiple trait animal model, as described by Tier and Graser (1991). In this method unique diagonal blocks of random effects are stored in memory only once and addressed by pointers when processing the permanent environment or animal effect. Animal effects were solved via Second Order Jacobi iteration with a relaxation factor of .7. All other effects were solved by the Gauss-Seidel method. The relative difference between consecutive solutions prior to relaxation was used to monitor convergence.

RESULTS AND DISCUSSION

Memory requirements were found to be no limitation for application of a multiple lactation test-day model to the national datasets of Canada and Germany. Although model TD2 for the German data set consisted of 14,704,144 equations, and 22 mio test day records were processed three times in every round of iteration, less than 200 MBytes of memory were required. Due to the implicit representation of the mixed model equations, test-day models TD1 and TD2 required only slightly more memory than lactation average models. One round of iteration took about 12 minutes CPU time on a HP 9000-891 workstation for model TD2 for the German data set. Processing time decreased by 25% if all test-day-records were stored in memory. Each record required 18 bytes of memory, thus overall demand of random access memory increased by about 400 MB for dataset 2. In this data set, 3104 different diagonal blocks for permanent environment effects on 1.4 mio cows were found. Consideration of inbreeding for construction of the A-1 increased number of different elements in A-1, but still only a total of 75005 different diagonal blocks for animals were found, although the complete dataset comprised of nearly 2.4 mio animals.

From detailed studies on dataset 1 (Table 3) it was concluded that 250 to 300 rounds of iteration were sufficient to reach stable solutions (relative difference, expressed on a log10 scale, reached a value of -4.0). A similar level of convergence was also reached for dataset 2 after 250 rounds.

The same mixed model equations (dataset 1) were also solved using solutions from a simulated previous (6 mo earlier) genetic evaluation run as starting values. The same level of convergence, which was reached after 300 rounds without starting values, was now reached after 150 rounds. Benefits from the use of starting values will likely be larger when more frequent evaluations (and therefore less new observations) are computed.

In model TD1, records from different lactations on the same sample day were fitted in different herd-test-date groups. For third lactation herd-test-date levels, subcell size becomes small in small herds and, therefore, the effective contribution of these records to genetic evaluation is small when herd-test-date effects are considered as fixed. For dataset one, 14.3% of all third lactation records were grouped in herd-test-date levels of size 1 or 2. Subcell sizes for herd-test-date levels for lactation 2 were larger, but still 6.1% of all second lactation records were grouped in herd-test-date levels of size 1 or 2. Subcell sizes for the same herd-test-date levels of size 1 or 2. Assigning records from second and third lactation to the same herd-test-date group, as in model TD2, resolved this problem, giving only 2.9% (3.4% for the German dataset) of second and third lactation records in herd-test-date levels of size 1 or 2. Differences between EBV from models TD1 and TD2 were extremely small for the combined EBV over all three lactations and therefore in further tables only results from TD2 will be used.

Table 4 shows correlations of combined EBV for SCS and protein yield from model TD2 with EBV from the three lactation average models for Canadian data. For SCS, differences between EBV for bulls from test-day models and lactation average models diminished with increasing numbers of daughters (Table 4). Lowest correlations were found with the first lactation model ST;

even with more than 100 daughters, correlations were as low as .86. Repeatability models REP1 and REP2 showed better agreement but correlations for bulls with 26 to 50 daughters, which might be characterized as young test bulls with their first crop of daughters, were .95. Correlations between cow EBV were in a range similar to correlations for bulls with a low number of daughters. For protein yield correlations were in general smaller than for SCS from this study and also lower than reported by Swalve (1995) for a study with 1st lactation German production data. It has to be investigated if differences in the present study are due to the difference in heritabilities used in the test day models (.25, .23, and .20 for lactations 1 to 3) compared to a heritability of .33 for lactation protein yield or if in general larger differences between test day models and lactation models can be found for multiple lactation TD models applied to production traits compared to SCS.

For production traits, most attention is paid to the ranking of animals with the most desirable (= highest) EBV. For SCS, focus would be mostly on avoiding bulls with high EBV for SCS, which would be undesirable. Table 5 displays the agreement of the ranking of bulls with highest EBV for SCS and protein yield from test day model TD2 to the ranking based on EBV from lactation average models REP1, REP2, and ST. As already indicated by the correlation between EBV, differences in ranking were largest for bulls with a low number of daughters. Agreement of TD2 was better with the repeatability models REP1 and REP2 than with model ST. With model ST, only 5 to 12 out of the 20 highest ranking bulls (for SCS) and 6 to 8 out of the 20 best bulls for protein yield from TD2 were ranked in the same group. In general differences between TD models and lactation average models were larger for protein yield than for SCS.

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TABLES

TABLE 1.	Number of test-day reco				
			2	3	Total
Dataset 1	Number of	3,081,676	1,633,087	790,909	5,505,672
CANADA	test day records	377,573	292,213	203,924	
	Number of herd- test-date levels	392,487	213,259	107,661	
D 41-41	Cows with records*	10,117,929	7,298,206	5,201,018	22,617,153
Dataset 2 GERMANY	test day records	1,576,243	1,595,682	1,595,682	
	Number of herd- test-date levels			629,283	
	Cows with records*	1,487,279	865,284	029,285	

mber of test-day records, cows, and levels of herd-test-date effects.

* Number of cows for lactation 2 and 3 are subsets of cows in lactation 1

	Vari	ances by lact	tation	Covariances between pairs of lactations			
Random effect	1	2	3	1:2	1:3	2:3	
CANADA <u>SCS</u> Animal additive genetic	.202	.222	.319	.187	.200	.251	
Permanent environment	.921	1.001	1.225	.276	.202	.504	
Residual	1.054	1.213	1.269	0	0	0	
Protein Yield, (kg * 100)							
Animal additive genetic	.398	.676	.667	.474	.443	.634	
Permanent environment	.569	1.098	1.272	.417	.324	.667	
Residual	.635	1.173	1.365	0	0	0	
GERMANY <u>SCS</u>							
Animal additive genetic	.154	.270	.323	.183	.189	.286	
Permanent environment	.737	.695	.764	.248	.209	.327	
Residual	1.053	1.098	1.159	0	0	0	

TABLE 2. Variance components used for genetic eval. of SCS with models TD1 and TD2.

 TABLE 3.
 Convergence criteria at various rounds of iteration for test-day model TD1 for genetic evaluation of SCS (dataset 1, correlations and deviations only for cows with observations).

Log ₁₀ of relative			Correlation of		Absolute deviation of EBV with EBV at round 500					
diffe (EBV with EBV at round 500		Mean		Maximum			
Iteration	Perm. envir.	EBV	Lact. 1	Lact. 3	Lact. 1	Lact. 3	Lact. 1	Lact. 3		
20	-2.398	-1.422		•••		***	••••			
50	-3.138	-2.897	.995	.994	.020	.051	.148	.315		
100	-3.742	-3.442	.999	.999	.009	.025	.062	.166		
150	-4.061	-3.746	•••		•••	•••		• • •		
200	-4.328	-4.000	1.000	1.000	.003	.009	.022	.080		
250	-4.536	-4.205					•••			
300	-4.681	-4.365	1.000	1.000	.001	.005	.009	.045		
400	-4.831	-4.559	1.000	1.000	.000	.002	.003	.019		
500	-4.896	-4.697	-	-						

		Number	Estimated breeding values from model						
EBV TD2		of animals	REP1		REP2		ST		
			SCS	PY	SCS	PY	SCS	PY	
Bulls	No. of daughters								
	5 -25	2598	.931	.728	.890	.712	.745	.688	
	26 - 50	1137	.954	.780	.917	.771	.814	.744	
	51 - 100	459	.968	.814	.940	. 8 02	.807	.757	
	> 100	320	.972	.839	.961	.813	.858	.766	
Cows w	vith observations	380,178	.957	.784	.940	.761	.815	.740	

TABLE 4.Correlations between estimated breeding values (dataset 1) from test-day model TD2 (combined
EBV for lactations 1, 2, and 3) and EBV from three lactation average models (REP1, REP2,
and ST) for SCS and protein yield (PY).

TABLE 5. Ranking of bulls with highest estimated breeding values (dataset 1) for SCS (most undesirable) and protein yield (PY) based on model TD2 (combined evaluation over lactations 1, 2, and 3) compared to the ranking based on estimated breeding values from lactation average models.

Number of	Number of	Ranking for TD2 –	No. of bulls in common with TD2						
daughters included in EBV	bulls		REPI		REP2		ST		
			SCS	PY	SCS	PY	SCS	PY	
5 - 25	2598	Top 20 Top 50 Top 100	14 34 70	11 30 60	11 30 60	9 17 47	5 20 44	6 18 40	
26 - 50	1137	Top 20 Top 50 Top 100	15 38 80	11 37 76	11 37 76	7 24 51	11 22 55	5 23 51	
51 - 100	459	Top 20 Top 50 Top 100	17 42 87	17 39 82	17 39 82	11 32 62	12 27 70	8 29 61	
> 100	320	Top 20 Top 50 Top 100	15 42 89	15 41 89	15 41 89	14 33 75	11 34 73	8 32 69	